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"Extraktion quantifizierbarer Information aus komplexen Systemen"

A Multilevel Monte Carlo Algorithm for Lévy Driven Stochastic Differential Equations

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A multilevel Monte Carlo algorithm for Lévy driven stochastic differential equations

by

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Summary. This article introduces and analyzes multilevel Monte Carlo schemes for the evaluation of the expectation $\mathbb{E}[f(Y)]$, where $Y = (Y_t)_{t \in [0,1]}$ is a solution of a stochastic differential equation driven by a Lévy process. Upper bounds are provided for the worst case error over the class of all measurable real functions f that are Lipschitz continuous with respect to supremum norm. Moreover, the worst case errors are compared with the computational cost of the corresponding algorithms. If the driving Lévy process has Blumenthal-Getoor index β , we obtain as dominant term in the upper estimate $n^{-(\frac{1}{\beta \vee 1} - \frac{1}{2})}$, when the computational cost n tends to infinity.

Keywords. Multilevel Monte Carlo, numerical integration, quadrature, Lévy-driven stochastic differential equation.

2000 Mathematics Subject Classification. Primary 60G51, Secondary 60H10, 60J75

1 Introduction

In this article, we analyze numerical schemes for the evaluation of

$$S(f) := \mathbb{E}[f(Y)],$$

where $Y = (Y_t)_{t \in [0,1]}$ is a solution to a multivariate stochastic differential equation driven by a multidimensional Lévy process, and f is a Borel measurable mapping from the Skorokhod space of \mathbb{R}^{d_Y} -valued functions over the time interval [0,1] into the real numbers that is Lipschitz continuous with respect to the *supremum* norm.

This is a classical problem which appears for instance in finance, where Y models the risk neutral stock price and f denotes the payoff of a (possibly path dependent) option, and in the past several concepts have been employed for dealing with it. We refer in particular to [PT97], [Rub03], and [JKMP05] for an analysis of the Euler scheme for Lévy-driven SDEs.

Recently, Giles [Gil08b] introduced the so called *multilevel Monte Carlo method* in the context of stochastic differential equations, and this turned out to be very efficient when Y is a continuous diffusion. Indeed, it can even be shown that it is -in some sense- optimal [CDMR08], see also [Gil08a, Avi09] for further recent results and [MR09] for a survey and further references. In this article, we analyze a multilevel Monte Carlo algorithm for Lévy driven stochastic differential equations. We use approximations that simulate large jumps of the Lévy process and neglect small ones, while the Wiener process is treated as in the Euler scheme. We control the worst case error over the class of all Lipschitz functionals with Lipschitz constant one. Moreover, we relate the approximation error to the computational cost of the algorithm. For Lévy processes with small Blumenthal-Getoor index, we find the same asymptotics as obtained in Giles [Gil08b] in the classical setting, i.e., the order of the error in the computational time n is $n^{-1/2}(\log n)^{3/2}$. For large Blumenthal-Getoor index the asymptotics are dominated by the Lévy process, where the critical index is one.

1.1 Notation and results

Let us now introduce the main notation. We denote by $|\cdot|$ the Euclidean norm for vectors as well as the Frobenius norm for matrices. For h > 0, we put $B_h = \{x \in \mathbb{R}^{d_X} : |x| < h\}$. Moreover, we let D[0, 1] denote the space of càdlàg functions from [0, 1] to \mathbb{R}^{d_Y} , where \mathbb{R}^{d_Y} is the state space of Y to be defined below. The space D[0, 1] is endowed with the σ -field induced by the projections on the marginals (or, equivalently, the Borel- σ -field for the Skorokhod topology). Often, we consider supremum norm on D[0, 1] and we denote $||f|| = \sup_{t \in [0,1]} |f(t)|$ for $f \in D[0, 1]$. Furthermore, the set of Borel measurable functions $f : D[0, 1] \to \mathbb{R}$, that are Lipschitz continuous with Lipschitz constant one with respect to the supremum norm is denoted by Lip(1). In general, we write $f \sim g$ iff $\lim f/g = 1$, while $f \leq g$ stands for $\limsup f/g \leq 1$. Finally, $f \approx g$ means $0 < \liminf f/g \leq \limsup f/g < \infty$, and $f \preceq g$ means $\limsup f/g < \infty$.

In the following, $X = (X_t)_{t\geq 0}$ denotes a d_X -dimensional L^2 -Lévy process with Lévy measure ν , drift parameter b and Gaussian covariance matrix $\Sigma\Sigma^*$, see Section 2. The parameters ν , $\Sigma\Sigma^*$ and b uniquely determine the distribution of the Lévy process on D[0, 1].

We consider the stochastic integral equation

$$Y_t = y_0 + \int_0^t a(Y_{s-}) \, \mathrm{d}X_s \tag{1}$$

with deterministic initial value $y_0 \in \mathbb{R}^{d_Y}$. We impose a standard Lipschitz assumption on the function a, which implies, in particular, existence and strong uniqueness of the solution:

Assumption. For a fixed $K < \infty$, the function $a : \mathbb{R}^{d_Y} \to \mathbb{R}^{d_Y \times d_X}$ satisfies

$$|a(y) - a(y')| \le K|y - y'|$$

for all $y, y' \in \mathbb{R}^{d_Y}$. Furthermore, we have

$$|a(y_0)| \le K, \ 0 < \int |x|^2 \nu(\mathrm{d}x) \le K^2, \ |\Sigma| \le K \text{ and } |b| \le K.$$

For a general account on Lévy processes we refer the reader to the books by [Ber98] and [Sat99]. Moreover, concerning stochastic analysis, we refer the reader to the books by Protter [Pro05] and Applebaum [App04].

We consider a class \mathcal{A} of multilevel Monte Carlo algorithms together with a cost function cost : $\mathcal{A} \to [0, \infty)$ that are introduced explicitly in Section 4. For each algorithm $\hat{S} \in \mathcal{A}$, we denote by $\hat{S}(f)$ a real-valued random variable representing the random output of the algorithm when applied to a given measurable function $f : D[0, 1] \to \mathbb{R}$. We work in the real number model of computation, which means that we assume that arithmetic operations with real numbers and comparisons can be done in one time unit. Our cost function represents the runtime of the algorithm reasonably well when supposing that

- one can sample from the distribution $\nu|_{B_h^c}/\nu(B_h^c)$ for sufficiently small h > 0 and the uniform distribution on [0, 1] in constant time,
- one can evaluate a at any point $y \in \mathbb{R}^{d_Y}$ in constant time, and
- *f* can be evaluated for piecewise constant functions in less than a constant multiple of its breakpoints plus one time units.

Indeed, in that case, the average runtime to evaluate $\widehat{S}(f)$ is less than a constant multiple of $\operatorname{cost}(\widehat{S})$. We denote the *worst case error* of an algorithm $\widehat{S} \in \mathcal{A}$ over all functions $f \in \operatorname{Lip}(1)$ by

$$e^{2}(\widehat{S}) = \sup_{f \in \operatorname{Lip}(1)} \mathbb{E}[|S(f) - \widehat{S}(f)|^{2}].$$
(2)

Finally, we analyze the minimum of $e(\widehat{S})$ over the class of all algorithms $\widehat{S} \in \mathcal{A}$ with cost bounded by n,

$$e_{\mathcal{A}}(n) = \inf_{\substack{\widehat{S} \in \mathcal{A}:\\ \cot(\widehat{S}) \le n}} e(\widehat{S}) \qquad , n \ge 1.$$

Our main findings are summarized in the following theorem. Algorithms that satisfy these error bounds are provided in Section 4.

Theorem 1. Let $g: (0,\infty) \to (0,\infty)$ denote a decreasing and invertible function such that

$$\int \frac{|x|^2}{h^2} \wedge 1\,\nu(\mathrm{d}x) \le g(h) \quad \text{for all } h > 0.$$

(i) If the driving process X has no Brownian component, i.e., $\Sigma = 0$, and if there exists $\gamma > 0$ such that

$$g(h) \precsim \frac{1}{h(\log 1/h)^{1+\gamma}} \tag{3}$$

as $h \downarrow 0$, then

$$e_{\mathcal{A}}(n) \precsim \frac{1}{\sqrt{n}}.$$

(ii) If there exists $\gamma \geq 1/2$ such that

$$g(h) \precsim \frac{(\log 1/h)^{\gamma}}{h},$$

as $h \downarrow 0$, then

$$e_{\mathcal{A}}(n) \precsim \frac{1}{\sqrt{n}} (\log n)^{\gamma+1}.$$

(iii) If there exists $\gamma > 1$ with

$$g(\frac{\gamma}{2}h) \ge 2g(h) \tag{4}$$

for all sufficiently small h > 0, then

$$e_{\mathcal{A}}(n) \precsim \sqrt{n} g^{-1}(n).$$

Remark 1. Part (ii) and (iii) deal with stochastic differential equations with a Brownian component in the driving process X, i.e. $\Sigma \neq 0$, see Section 2. Essentially part (ii) and (iii) cover all reasonable cases with $\Sigma \neq 0$. When $\gamma = 1/2$ in (ii), the asymptotics are the same as in the classical diffusion setting analyzed in [Gil08b]. Similarly, as in our proof one can also treat different terms of lower order instead of log. Certainly, it also makes sense to consider $\gamma < 1/2$, when $\Sigma = 0$. The computations are similar and therefore omitted. Part (iii) covers, in particular, all cases where g is regularly varying at 0 with exponent strictly smaller than -1. In this case it is possible to choose $g(h) = \int \frac{|x|^2}{h^2} \wedge 1 \nu(dx)$.

In terms of the Blumenthal-Getoor index

$$\beta := \inf \left\{ p > 0 : \int_{B_1} |x|^p \, \nu(\mathrm{d}x) < \infty \right\} \in [0, 2]$$

we get the following corollary.

Corollary 1.

$$\sup\{\gamma \ge 0: e_{\mathcal{A}}(n) \precsim n^{-\gamma}\} \ge \left(\frac{1}{\beta} - \frac{1}{2}\right) \land \frac{1}{2}$$

1.2 Examples

We now apply our results of Theorem 1 to some common examples of driving Lévy processes.

- Jump diffusion processes are Brownian motions interlaced with a compound Poisson process. In finance, special cases of these are used, e.g., in the Kou or the Merton model, to model the log price process. The compound Poisson process has finite Lévy measure and thus the integral $\int \frac{|x|^2}{h^2} \wedge 1 \nu(dx)$ is bounded by a constant. Hence, the tradeoff between cost and error is the same as in the pure Gaussian case, i.e. $e_{\mathcal{A}}(n) \preceq n^{-1/2} (\log n)^{3/2}$.
- Since α -stable processes only satisfy our assumption of second moments in the special case of Brownian Motion ($\alpha = 2$), we consider processes where the jumps have α -stable behaviour around 0 and are truncated at some given size u > 0. The Lévy measure then has Lebesgue-density

$$\nu(x) = \frac{c_1}{x^{1+\alpha}} \mathbf{1}_{(0,u)}(x) + \frac{c_2}{|x|^{1+\alpha}} \mathbf{1}_{(-u,0)}(x)$$

and the complexity functional for small h > 0 is bounded by

$$\int \frac{|x|^2}{h^2} \wedge 1 \ \nu(dx) \le Ch^{-\alpha} = g(h)$$

with a constant C > 0 depending on c_1, c_2, u and α .

For $\alpha < 1$ part (i) can be used because (3) is fulfilled for $\gamma > 0$ and $h \downarrow 0$. The asymptotic error for cost n thus behaves like $e_{\mathcal{A}}(n) \preceq n^{-\frac{1}{2}}$.

In the case with $\alpha > 1$ part (iii) can be used with $\gamma = 2^{\frac{\alpha-1}{\alpha}} > 1$ and the error bound for cost *n* fulfills $e_{\mathcal{A}}(n) \preceq n^{-(\frac{1}{\alpha} - \frac{1}{2})}$.

The case $\alpha = 1$ can be treated via (ii). However, this is suboptimal since in this case the error without Gaussian term is smaller than the one with Gaussian term. Similar computations as in the proof of Theorem 1 show that for the truncated 1-stable Lévy process one has $e_{\mathcal{A}}(n) \preceq n^{-\frac{1}{2}} \log n$.

2 Basic Facts

The distribution of the driving Lévy process X is characterized by the parameters $b \in \mathbb{R}^{d_X}, \Sigma \in \mathbb{R}^{d_X \times d_X}$ and ν , which is a Borel measure on $\mathbb{R}^{d_X} \setminus \{0\}$ and satisfies

$$0 < \int |x|^2 \nu(\mathrm{d}x) < \infty.$$

We sketch a construction of X with a view towards simulation in the multilevel setting. This construction is based on the L^2 -approximation of Lévy processes as it is presented in [Pro05] and [App04].

Consider a stochastic process $(N(t, A))_{t \ge 0, A \in \mathfrak{B}(\mathbb{R}^{d_X \setminus \{0\}})}$ on some probability space $(\Omega, \mathfrak{A}, P)$ with the following properties. For every $\omega \in \Omega$ the mapping $[0, t] \times A \mapsto N(t, A)(\omega)$ induces a σ -finite counting measure on $\mathfrak{B}(\mathbb{R}_+ \times (\mathbb{R}^{d_X} \setminus \{0\}))$. For every $A \in \mathfrak{B}(\mathbb{R}^{d_X} \setminus \{0\})$ that is bounded away from zero the process $(N(t, A))_{t \ge 0}$ is a Poisson process with intensity $\nu(A)$. For pairwise disjoint sets $A_1, \ldots, A_r \in \mathfrak{B}(\mathbb{R}^{d_X} \setminus \{0\})$ the stochastic processes $(N(t, A_1))_{t \ge 0}, \ldots, (N(t, A_r))_{t \ge 0}$ are independent.

Then N(t, .) *P*-a.s. defines a finite measure on B_h^c with values in \mathbb{N}_0 . The integral $\int_{B_h^c} x N(t, dx)$ thus is a random finite sum, which gives rise to a compound Poisson process. More specifically put $\lambda^{(h)} = \nu(B_h^c) < \infty$, which satisfies $\lambda^{(h)} > 0$ for sufficiently small h > 0. In this case $\mu^{(h)} = \nu|_{B_h^c}/\lambda^{(h)}$ defines a probability measure on $\mathbb{R}^{d_X} \setminus \{0\}$ such that

$$\int_{B_h^c} x N(t, \mathrm{d}x) \stackrel{d}{=} \sum_{i=1}^{N_t} \xi_i,\tag{5}$$

where $\stackrel{d}{=}$ denotes equality in distribution, $(N_t)_{t\geq 0}$ is a Poisson process with intensity $\lambda^{(h)}$ and $(\xi_i)_{i\in\mathbb{N}}$ is an i.i.d. sequence of random variables with distribution $\mu^{(h)}$ and independent of $(N_t)_{t\geq 0}$. Its expectation calculates to $\mathbb{E}\left[\int_{B_h^c} x N(t, \mathrm{d}x)\right] = F_0(h)t$, where we set

$$F_0(h) = \int_{B_h^c} x \,\nu(\mathrm{d}x).$$

The compensated process $L^{(h)} = (L_t^{(h)})_{t \ge 0}$, given by

$$L_t^{(h)} = \int_{B_h^c} x N(t, \mathrm{d}x) - tF_0(h),$$
(6)

is an L^2 -martingale, and the same holds true for its L^2 -limit $L = (L_t)_{t \ge 0} = \lim_{h \downarrow 0} L^{(h)}$, see, e.g., Applebaum [App04].

With W denoting a d_X -dimensional Brownian motion independent of L, we define the Lévy process X by

$$X_t = \Sigma W_t + L_t + bt. \tag{7}$$

We add that the Lévy-Itô-decomposition guarantees that every L^2 -Lévy process has a representation (7).

3 A Coupled Euler Scheme

The multilevel Monte Carlo algorithm introduced in Section 4 is based on a hierarchy of coupled Euler schemes for the approximation of the solution process Y of the SDE (1). In every single Euler scheme we approximate X in the following way. At first we neglect all the jumps with size smaller than h. The jumps of size at least h induce a random time discretization, which, because of the Brownian component, is refined so that the step sizes are at most ε . Finally W as well as the drift component are approximated by piecewise constant functions with respect to the refined time discretization. In this way, we get an approximation $\hat{X}^{(h,\epsilon)}$ of X. The multilevel approach requires simulation of the joint distribution of $\hat{X}^{(h,\epsilon)}$ and $\hat{X}^{(h',\epsilon')}$ for different values of h' > h > 0 and $\varepsilon' > \varepsilon > 0$. More precisely, we proceed as follows.

For any càdlàg process L we denote by $\Delta L_t = L_t - \lim_{s \nearrow t} L_s$ the jump-discontinuity at time t. The jump times of $L^{(h)}$ are then given by $T_0^{(h)} = 0$ and

$$T_k^{(h)} = \inf\{t > T_{k-1}^{(h)} : \Delta L_t^{(h)} \neq 0\}$$

for $k \geq 1$. The time differences $T_k^{(h)} - T_{k-1}^{(h)}$ form an i.i.d. sequence of random variables that are exponentially distributed with parameter $\lambda^{(h)}$. Furthermore, this sequence is independent of the sequence of jump heights $\Delta L_{T_k^{(h)}}^{(h)}$, which is i.i.d. with distribution $\mu^{(h)}$, and on every interval

 $[T_k^{(h)}, T_{k+1}^{(h)}]$ the process $L^{(h)}$ is linear with slope $-F_0(h)$. See (5) and (6).

The processes $\Delta L^{(h')}$ and $\Delta (L^{(h)} - L^{(h')})$ are independent with values in $\{0\} \cup B_{h'}^c$ and $\{0\} \cup B_h^c \setminus B_{h'}^c$, respectively, and therefore the jumps of the process $L^{(h')}$ can be obtained from those of $L^{(h)}$ by

$$\Delta L_t^{(h')} = \Delta L_t^{(h)} \cdot \mathbf{1}_{\{|\Delta L_t^{(h)}| > h'\}}.$$

We conclude that the simulation of the joint distribution of $(L^{(h)}, L^{(h')})$ only requires samples from the jump times and jump heights $T_k^{(h)}$ and $\Delta L_{T_k^{(h)}}^{(h)}$, respectively, which amounts to sampling from $\mu^{(h)}$ and from an exponential distribution.

Because of the Brownian component we refine the time discretization by $T_0^{(h,\varepsilon)} = 0$ and

$$T_j^{(h,\varepsilon)} = \inf\{T_k^{(h)} > T_{j-1}^{(h,\varepsilon)} : k \in \mathbb{N}\} \land (T_{j-1}^{(h,\varepsilon)} + \varepsilon)$$

$$\tag{8}$$

for $j \ge 1$. Summarizing, X is approximated at the discretization times $T_j = T_j^{(h,\varepsilon)}$ by $\hat{X}_0^{(h,\varepsilon)} = 0$ and

$$\hat{X}_{T_j}^{(h,\epsilon)} = \hat{X}_{T_{j-1}}^{(h,\epsilon)} + \Sigma(W_{T_j} - W_{T_{j-1}}) + \Delta L_{T_j}^{(h)} + (b - F_0(h))(T_j - T_{j-1})$$

for $j \geq 1$. Observe that

$$\hat{X}_{T_j}^{(h,\epsilon)} = \Sigma W_{T_j} + L_{T_j}^{(h)} + bT_j.$$

To simulate the Brownian components of the coupled processes $(\hat{X}^{(h,\epsilon)}, \hat{X}^{(h',\epsilon')})$, we refine the sequence of jump times $T_k^{(h)}$ to get $(T_j^{(h,\epsilon)})_{j\in\mathbb{N}_0}$ and $(T_j^{(h',\epsilon')})_{j\in\mathbb{N}_0}$, respectively. Since W and L are independent, the process W is easily simulated at all times $(T_j^{(h,\epsilon)})_{j\in\mathbb{N}_0}$ and $(T_j^{(h',\epsilon')})_{j\in\mathbb{N}_0}$ that are in [0, 1] by sampling from a normal distribution.

For the SDE (1) the Euler scheme with the driving process $(\Sigma W_t + L_t^{(h)} + bt)_{t\geq 0}$ and the random time discretization $(T_j)_{j\in\mathbb{N}_0} = (T_j^{(h,\epsilon)})_{j\in\mathbb{N}_0}$ is defined by $\hat{Y}_0^{(h,\epsilon)} = y_0$ and

$$\hat{Y}_{T_{j}}^{(h,\epsilon)} = \hat{Y}_{T_{j-1}}^{(h,\epsilon)} + a(\hat{Y}_{T_{j-1}}^{(h,\epsilon)})(\hat{X}_{T_{j}}^{(h,\epsilon)} - \hat{X}_{T_{j-1}}^{(h,\epsilon)})$$

$$\tag{9}$$

for $j \geq 1$. Furthermore $\hat{Y}_t^{(h,\epsilon)} = \hat{Y}_{T_j}^{(h,\epsilon)}$ for $t \in [T_j, T_{j+1})$. In the multilevel approach the solution process Y of the SDE (1) is approximated via coupled Euler schemes $(\hat{Y}^{(h,\epsilon)}, \hat{Y}^{(h',\epsilon')})$, which are obtained by applying the Euler scheme (9) to the coupled driving processes $\hat{X}^{(h,\epsilon)}$ and $\hat{X}^{(h',\epsilon')}$ with their random discretization times $T_j^{(h,\epsilon)}$ and $T_j^{(h',\epsilon')}$, respectively.

4 The Multilevel Monte Carlo Algorithm

We fix two positive and decreasing sequences $(\varepsilon_k)_{k\in\mathbb{N}}$ and $(h_k)_{k\in\mathbb{N}}$, and we put $Y_t^{(k)} = \hat{Y}_t^{(h_k,\epsilon_k)}$. For technical reasons we define $Y_t^{(k)}$ for all $t \ge 0$, although we are typically only interested in $Y^{(k)} := (Y_t^{(k)})_{t\in[0,1]}$. For $m \in \mathbb{N}$ and a given measurable function $f : D[0,1] \to \mathbb{R}$ with $f(Y^{(k)})$ being integrable for $k = 1, \ldots, m$, we write $\mathbb{E}[f(Y^{(m)})]$ as telescoping sum

$$\mathbb{E}[f(Y^{(m)})] = \mathbb{E}[f(Y^{(1)})] + \sum_{k=2}^{m} \mathbb{E}[f(Y^{(k)}) - f(Y^{(k-1)})].$$

In the multilevel approach each expectation on the right-hand side is approximated seperately by means of independent classical Monte Carlo approximations. For k = 2, ..., m we denote by n_k the number of replications for the approximation of $\mathbb{E}[f(Y^{(k)}) - f(Y^{(k-1)})]$ and by n_1 the number of replications for the approximation of $\mathbb{E}[f(Y^{(1)})]$. For $(Z_{j,1}^{(k)}, Z_{j,2}^{(k)})_{j=1,...,n_k}$ being i.i.d. copies of the coupled Euler scheme $(Y^{(k-1)}, Y^{(k)})$ for k = 2, ..., m and $(Z_j^{(1)})_{j=1,...,n_1}$ being i.i.d. copies of $Y^{(1)}$, the corresponding multilevel Monte Carlo algorithm is given by

$$\widehat{S}(f) = \frac{1}{n_1} \sum_{j=1}^{n_1} f(Z_j^{(1)}) + \sum_{k=2}^m \frac{1}{n_k} \sum_{j=1}^{n_k} \left[f(Z_{j,2}^{(k)}) - f(Z_{j,1}^{(k)}) \right].$$

The algorithm \widehat{S} is uniquely described by the parameters m and $(n_k, h_k, \varepsilon_k)_{k=1,...,m}$ so that we formally identify the algorithm \widehat{S} with these parameters. We denote by \mathcal{A} the set of all algorithms \widehat{S} that are of this form.

The error of the algorithm

For measurable functions $f: D[0,1] \to \mathbb{R}$ with $f(Y), f(Y^{(1)}), \ldots, f(Y^{(m)})$ being square integrable, the mean squared error of $\widehat{S}(f)$ calculates to

$$\mathbb{E}[|S(f) - \widehat{S}(f)|^2] = |\mathbb{E}[f(Y) - f(Y^{(m)})]|^2 + \operatorname{var}(\widehat{S}(f)).$$

If f is in Lip(1), then

$$\mathbb{E}[|S(f) - \widehat{S}(f)|^2] \le \mathbb{E}||Y - Y^{(m)}||^2 + \sum_{k=2}^m \frac{1}{n_k} \mathbb{E}||Y^{(k)} - Y^{(k-1)}||^2 + \frac{1}{n_1} \mathbb{E}||Y^{(1)} - y_0||^2.$$
(10)

In particular, the upper bound does not depend on the choice of f. Hence (10) remains valid for the worst case error $e^2(\widehat{S})$ as defined in (2).

The cost of the algorithm

For a piecewise constant \mathbb{R}^{d_Y} -valued function $y = (y_t)_{t \in [0,1]}$, we denote by $\Upsilon(y)$ its number of breakpoints. Then the cost of the algorithm \widehat{S} is defined by the function

$$\operatorname{cost}(\widehat{S}) = \sum_{k=1}^{m} n_k \mathbb{E}[\Upsilon(Y^{(k)})].$$
(11)

Note that under the assumptions quoted in Section 1 and if $\varepsilon_1 \leq 1$, the average runtime of the algorithm \widehat{S} is indeed less than a constant multiple of $\operatorname{cost}(\widehat{S})$.

The choice of the parameters

The algorithm \widehat{S} is completely determined by the parameters m and $(n_k, \varepsilon_k, h_k)_{k=1,...,m}$ and we now give the parameters which achieve the error estimates provided in Theorem 1 and which correspond to algorithms with cost of order at most n. Recall that Theorem 1 depends on an invertible and decreasing function $g: (0, \infty) \to (0, \infty)$ satisfying

$$\int \frac{|x|^2}{h^2} \wedge 1\,\nu(\mathrm{d}x) \le g(h) \text{ for all } h > 0,$$

and we set, for $k \in \mathbb{N}$,

$$\varepsilon_k = 2^{-k}$$
 and $h_k = g^{-1}(2^k)$.

We choose the remaining parameters by

 $m = \inf\{k \in \mathbb{N}: Ch_k < 1\} - 1 \quad \text{and} \quad n_k = \lfloor Ch_{k-1} \rfloor \quad \text{for } k = 1, \dots, m, \quad (12)$

where

(i)
$$C = n$$
, (ii) $C = \frac{n}{(\log n)^{\gamma+1}}$, and (iii) $C = 1/g^{-1}(n)$

in the respective cases. Here we need to assume additionally that g is such that $h^{-2/3} \preceq g(h)$ in case (i) and $h^{-1}\sqrt{\log 1/h} \preceq g(h)$ in case (ii). These parameters optimize (up to constant multiples) the error estimate induced by equation (10) together with Theorem 2 below.

5 Proofs

Proofing the main result requires asymptotic error bounds for the strong approximation of Y by $\hat{Y}^{(h,\epsilon)}$ for given $\varepsilon, h > 0$. Therefore we define for h > 0

$$F(h) = \int_{B_h} |x|^2 \,\nu(\mathrm{d}x)$$

to approximate the error that originates from the neglect of the jumps smaller than h. This is a reasonable choice as we consider an L^2 average of the supremum error and as the following theorem reveals.

Theorem 2. Under Assumption 1.1, there exists a constant κ depending only on K such that for any $\varepsilon \in (0,1]$ and h > 0 with $\nu(B_h^c) \leq \frac{1}{\varepsilon}$, one has

$$\mathbb{E}\left[\sup_{t\in[0,1]}|Y_t-\hat{Y}_t^{(h,\epsilon)}|^2\right] \le \kappa \left(\varepsilon \log(e/\varepsilon) + F(h)\right)$$

in the general case and

$$\mathbb{E}\Big[\sup_{t\in[0,1]}|Y_t-\hat{Y}_t^{(h,\epsilon)}|^2\Big] \le \kappa \big[F(h)+|b-F_0(h)|^2\varepsilon^2\big]$$

in the case without Wiener process, i.e. $\Sigma = 0$.

Remark 2. A similar Euler scheme is analyzed in [Rub03]. There it is shown that the appropriately scaled error process (the discrepancy between approximative and real solution) converges in distribution to a stochastic process. In the cases where this limit theorem is applicable, it is straight-forward to verify that the estimate provided in Theorem 2 is of the right order.

Remark 3. In the case without Wiener process the term $|b - F_0(h)|^2 \varepsilon^2$ is typically of lower order than F(h), so that we have in most cases that

$$\mathbb{E}\left[\sup_{t\in[0,1]}|Y_t-\hat{Y}_t^{(h,\epsilon)}|^2\right]\lesssim\kappa\,F(h)$$

The proof of Theorem 2 is based on the analysis of an auxiliary process (\bar{Y}_t) : We decompose the Lévy martingale L into a sum of the Lévy martingale $L' = L^{(h)}$ constituted by the sum of compensated jumps of size bigger than h and the remaining part $L'' = (L_t - L'_t)_{t\geq 0}$. We denote $\bar{X} = (\Sigma W_t + L'_t + tb)_{t\geq 0}$, and let $\bar{Y} = (\bar{Y}_t)_{t\geq 0}$ be the solution to the integral equation

$$\bar{Y}_t = y_0 + \int_0^t a(\bar{Y}_{\iota(s-)}) \,\mathrm{d}\bar{X}_s,\tag{13}$$

where $\iota(t) = \sup[0, t] \cap \mathbb{T}$, and \mathbb{T} is the set of random times $(T_j)_{j \in \mathbb{Z}_+}$ defined by $T_j = T_j^{(h,\varepsilon)}$ as in (8) in Section 3.

Proposition 1. Under Assumption 1.1, there exists a constant κ depending only on K such that for any $\varepsilon \in (0, 1]$ and h > 0 with $\nu(B_h^c) \leq 1/\varepsilon$ we have

$$\mathbb{E}\left[\sup_{t\in[0,1]}|Y_t-\bar{Y}_t|^2\right] \le \kappa\left(\varepsilon+F(h)\right)$$

in the general case and

$$\mathbb{E}\left[\sup_{t\in[0,1]}|Y_t-\bar{Y}_t|^2\right] \le \kappa \left[F(h)+|b-F_0(h)|^2\varepsilon^2\right]$$

in the case without Wiener process, i.e. $\Sigma = 0$.

The proof of the proposition relies on the following lemma.

Lemma 1. Under Assumption 1.1, we have

$$\mathbb{E}\Big[\sup_{t\in[0,1]}|Y_t-y_0|^2\Big]\leq\kappa,$$

where κ is a finite constant depending on K only.

The proof of the lemma can be achieved along the standard argument for proving bounds for second moments. Indeed, the standard combination of Gronwall's lemma together with Doob's inequality yields the result.

Proof of Proposition 1. For $t \ge 0$, we consider $Z_t = Y_t - \bar{Y}_t$ and $Z'_t = Y_t - \bar{Y}_{\iota(t)}$. We fix a stopping time τ and let $z_{\tau}(t) = \mathbb{E}[\sup_{s \in [0, t \land \tau]} |Z_s|^2]$. To indicate that a process is stopped at time τ we put τ in its superscript. The main task of the proof is to establish an estimate of the form

$$z_{\tau}(t) \le \alpha_1 \int_0^t z_{\tau}(s) \,\mathrm{d}s + \alpha_2$$

with values $\alpha_1, \alpha_2 > 0$ that do not depend on the choice of τ . Then by using a localizing sequence of stopping times (τ_n) with finite $z_{\tau_n}(1)$, we deduce from Gronwall's inequality that

$$\mathbb{E}[\sup_{s\in[0,1]}|Y_s-\bar{Y}_s|^2] = \lim_{n\to\infty} z_{\tau_n}(1) \le \alpha_2 \exp(\alpha_1).$$

We analyze

$$Z_{t} = \underbrace{\int_{0}^{t} (a(Y_{s-}) - a(\bar{Y}_{\iota(s-)})) \, \mathrm{d}(\Sigma W_{s} + L'_{s}) + \int_{0}^{t} a(Y_{s-}) \, \mathrm{d}L''_{s}}_{=:M_{t}} + \int_{0}^{t} (a(Y_{s-}) - a(\bar{Y}_{\iota(s-)})) \, b \, \mathrm{d}s$$
(14)

with $M = (M_t)_{t \ge 0}$ being a local L^2 -martingale. By Doob and Lemma 3, we get

$$\mathbb{E}\sup_{s\in[0,t\wedge\tau]}|M_s|^2 \le 4\,\mathbb{E}\Big[\int_0^{t\wedge\tau}|a(Y_{s-})-a(\bar{Y}_{\iota(s-)})|^2\,\mathrm{d}\langle\Sigma W+L'\rangle_s+\int_0^{t\wedge\tau}|a(Y_{s-})|^2\,\mathrm{d}\langle L''\rangle_s\Big],$$

where in general for a local L^2 -martingale A we set $\langle A \rangle_t = \sum_j \langle A^{(j)} \rangle_t$, where $\langle A^{(j)} \rangle$ denotes the predictable compensator of the classical bracket process of the *j*-th coordinate of A. Note that $d\langle \Sigma W + L' \rangle_t = (|\Sigma|^2 + \int_{B_h^c} |x|^2 \nu(dx)) dt \leq 2K^2 dt$ and similarly $d\langle L'' \rangle_t = F(h) dt$. Consequently, by Assumption 1.1 and Fubini's theorem, we get

$$\mathbb{E} \sup_{s \in [0, t \wedge \tau]} |M_s|^2 \le 8K^4 \int_0^t \mathbb{E}[\mathbb{1}_{\{s \le \tau\}} |Z_{s-}^{\prime \tau}|^2] \,\mathrm{d}s + 4K^2 F(h) \int_0^t \mathbb{E}[(|Y_{s-} - y_0| + 1)^2] \,\mathrm{d}s.$$

Conversely, by the Cauchy-Schwarz inequality and Fubini's theorem, one has for $t \in [0, 1]$ that

$$\mathbb{E}\left[\left|\int_{0}^{t\wedge\tau} (a(Y_{s-}) - a(\bar{Y}_{\iota(s-)})) \, b \, \mathrm{d}s\right|^{2}\right] \le K^{4} \int_{0}^{t} \mathbb{E}[\mathbb{1}_{\{s\le\tau\}} |Z_{s-}^{\prime\tau}|^{2}] \, \mathrm{d}s.$$

Using that for $a, b \in \mathbb{R}$, $(a + b)^2 \le 2a^2 + 2b^2$, we deduce with (14) that

$$\mathbb{E}\sup_{s\in[0,t\wedge\tau]}|Z_s|^2 \le 18K^4 \int_0^t \mathbb{E}[\mathbb{1}_{\{s\le\tau\}}|Z_{s-}^{\prime\tau}|^2]\,\mathrm{d}s + 8K^2F(h)\int_0^t \mathbb{E}[(|Y_{s-}-y_0|+1)^2]\,\mathrm{d}s.$$

Since $Z'_t = Z_t + \bar{Y}_t - \bar{Y}_{\iota(t)}$ we conclude that

$$\mathbb{E}\sup_{s\in[0,t\wedge\tau]}|Z_s|^2 \le 36K^4 \int_0^t \left[\mathbb{E}[|Z_{s-}^{\tau}|^2] + \mathbb{E}[\mathbb{1}_{\{s\leq\tau\}}|\bar{Y}_{s-} - \bar{Y}_{\iota(s-)}|^2]\right] \mathrm{d}s + 8K^2F(h)\int_0^t \mathbb{E}[(|Y_{s-} - y_0| + 1)^2] \mathrm{d}s$$

By Lemma 1, $\mathbb{E}[\sup_{s \in [0,1]}(|Y_s - y_0| + 1)^2]$ is bounded by a constant, and we get, for $t \in [0,1]$,

$$z_{\tau}(t) \le \kappa_1 \Big[\int_0^t \Big[z_{\tau}(s) + \mathbb{E}[\mathbb{1}_{\{s \le \tau\}} | \bar{Y}_{s-} - \bar{Y}_{\iota(s-)} |^2] \Big] \,\mathrm{d}s + F(h) \Big], \tag{15}$$

where κ_1 is a constant that depends only on K.

Next, we provide an appropriate estimate for $\mathbb{E}[\mathbb{1}_{\{s \leq \tau\}} | \bar{Y}_t - \bar{Y}_{\iota(t)} |^2]$. Since L' has no jumps on $(\iota(t), t)$ we have

$$\bar{Y}_t - \bar{Y}_{\iota(t)} = a(\bar{Y}_{\iota(t)}) \left[\Sigma(W_t - W_{\iota(t)}) + (b - F_0(h))(t - \iota(t)) \right]$$

so that

$$\mathbb{E}[\mathbb{1}_{\{t \le \tau\}} | \bar{Y}_t - \bar{Y}_{\iota(t)} |^2] \le 2K^2 \mathbb{E}[(|\bar{Y}_{\iota(t)}^{\tau} - y_0| + 1)^2] \left[|\Sigma|^2 \varepsilon + |b - F_0(h)|^2 \varepsilon^2 \right].$$

Here we used the independence of the Wiener process and the random times in \mathbb{T} . Next, we use that $|\bar{Y}_{\iota(t)} - y_0| \leq |Y_{\iota(t)} - y_0| + |Z_{\iota(t)}|$ to deduce that

$$\mathbb{E}[\mathbb{1}_{\{t \le \tau\}} | \bar{Y}_t - \bar{Y}_{\iota(t)} |^2] \le 4K^2 \big[\mathbb{E}[(|Y_{\iota(t)}^{\tau} - y_0| + 1)^2] + \mathbb{E}[|Z_{\iota(t)}^{\tau}|^2] \big] \big[|\Sigma|^2 \varepsilon + |b - F_0(h)|^2 \varepsilon^2 \big].$$

Recall that $\mathbb{E}[\sup_{s\in[0,1]}(|Y_s-y_0|+1)^2]$ is uniformly bounded. Moreover, by the Cauchy-Schwarz inequality, $|F_0(h)|^2 \leq K^2\nu(B_h^c) \leq K^2/\varepsilon$ so that the right bracket in the latter equation is uniformly bounded. Consequently, for $t \in [0,1]$,

$$\mathbb{E}[\mathbb{1}_{\{t \le \tau\}} | \bar{Y}_t - \bar{Y}_{\iota(t)} |^2] \le \kappa_2 \left[|\Sigma|^2 \varepsilon + |b - F_0(h)|^2 \varepsilon^2 + z_\tau(t) \right]$$

where κ_2 is an appropriate constant that depends only on K.

Inserting this estimate into (15) gives

$$z_{\tau}(t) \leq \kappa_3 \Big[\int_0^t z_{\tau}(s) \,\mathrm{d}s + |\Sigma|^2 \varepsilon + |b - F_0(h)|^2 \varepsilon^2 + F(h) \Big].$$

for a constant κ_3 that depends only on K. If $\Sigma = 0$, then the statement of the proposition follows from the Gronwall inequality. The general case is an immediate consequence of the estimates

$$|\Sigma|^2 \varepsilon \le K^2 \varepsilon$$
 and $|b - F_0(h)|^2 \varepsilon^2 \le 2K^2 (\varepsilon^2 + \varepsilon) \le 4K^2 \varepsilon$,

where we used again that $|F_0(h)|^2 \leq K^2/\varepsilon$.

For the proof of Theorem 2, we need a further lemma.

Lemma 2. Let $r \in \mathbb{N}$ and $(\mathcal{G}_j)_{j=0,1,\ldots,r}$ denote a filtration. Moreover, let, for $j = 0, \ldots, r-1$, U_j and V_j denote non-negative random variables such that U_j is \mathcal{G}_j -measurable, and V_j is \mathcal{G}_{j+1} -measurable and independent of \mathcal{G}_j . Then one has

$$\mathbb{E}\left[\max_{j=0,\dots,r-1}U_jV_j\right] \le \mathbb{E}\left[\max_{j=0,\dots,r-1}U_j\right] \cdot \mathbb{E}\left[\max_{j=0,\dots,r-1}V_j\right].$$

Proof. Without loss of generality we can and will assume that (U_j) is monotonically increasing. Otherwise, we can prove the result for (\tilde{U}_j) given by $\tilde{U}_j = \max_{k \leq j} U_k$ instead, and then deduce the result for the original sequence (U_j) .

We proceed by induction. For r = 1 the statement is trivial, since U_0 and V_0 are independent random variables. Next, we let $r \ge 1$ arbitrary and note that

$$\mathbb{E}[\max_{j=0,\dots,r} U_j V_j] = \mathbb{E}[\max_{j=1,\dots,r} U_j V_j] + \mathbb{E}[(U_0 V_0 - \max_{j=1,\dots,r} U_j V_j)^+].$$

Using the monotonicity of (U_j) , we get that

$$\mathbb{E}[(U_0V_0 - \max_{j=1,\dots,r} U_jV_j)^+ | \mathcal{G}_0] \le U_0 \mathbb{E}[(V_0 - \max_{j=1,\dots,r} V_j)^+ | \mathcal{G}_0] = U_0 \mathbb{E}[(V_0 - \max_{j=1,\dots,r} V_j)^+].$$

Next, we use the induction hypothesis for $\mathbb{E}[\max_{j=1,\dots,r} U_j V_j]$ to deduce that

$$\mathbb{E}[\max_{j=0,\dots,r} U_j V_j] \leq \mathbb{E}[U_r] \mathbb{E}[\max_{j=1,\dots,r} V_j] + \mathbb{E}[U_0] \mathbb{E}[(V_0 - \max_{j=1,\dots,r} V_j)^+]$$
$$\leq \mathbb{E}[U_r] \mathbb{E}[\max_{j=0,\dots,r} V_j].$$

Proof of Theorem 2. By Proposition 1, it remains to find an appropriate upper bound for $\mathbb{E}[\sup_{t \in [0,1]} |\bar{Y}_t - \hat{Y}_t|^2]$ where we denote $\hat{Y}_t = \hat{Y}_t^{(h,\epsilon)}$. First note that for all $j \in \mathbb{N}$, one has $\Delta L_{T_j} = \Delta L'_{T_j}$ and

$$L'_{T_{j+1}} = L'_{T_j} + \Delta L'_{T_{j+1}} - (T_{j+1} - T_j)F_0(h)$$

so that by definition, the processes (\bar{Y}_t) and (\hat{Y}_t) coincide almost surely for all times in \mathbb{T} (see (13) and (9)). Hence,

$$\bar{Y}_t - \hat{Y}_t = \bar{Y}_t - \bar{Y}_{\iota(t)} = \underbrace{a(\bar{Y}_{\iota(t)})\Sigma(W_t - W_{\iota(t)})}_{=:A_t} + \underbrace{a(\bar{Y}_{\iota(t)})(b - F_0(h))(t - \iota(t))}_{=:B_t}$$

Since two neighboring points in \mathbb{T} are at most ε units apart we get

$$\mathbb{E}\Big[\sup_{t\in[0,1]}|B_t|^2\Big] \le K^2 \,\mathbb{E}[(\|\bar{Y}-y_0\|+1)^2]|b-F_0(h)|^2\varepsilon^2.$$
(16)

It remains to analyze

$$\mathbb{E}\Big[\sup_{t\in[0,1]}|A_t|^2\Big] \le K^2|\Sigma|^2 \mathbb{E}\Big[\sup_{t\in[0,1]}(|\bar{Y}_{\iota(t)}-y_0|+1)^2|W_t-W_{\iota(t)}|^2\Big]$$

First suppose that (L'_t) is a *deterministic* piecewise linear càdlàg path. Then the times (T_j) are deterministic and we denote by r the minimal index j with $T_j \ge 1$. We estimate

$$\mathbb{E}\Big[\sup_{t\in[0,1]}|A_t|^2\Big] \le K^2 |\Sigma|^2 \mathbb{E}\Big[\sup_{j=0,\dots,r-1} \left((|\bar{Y}_{T_j}-y_0|+1)^2 \cdot \sup_{t\in[T_j,T_{j+1}\wedge 1)} |W_t-W_{T_j}|^2 \right) \Big].$$

Next, we apply Lemma 2 with $U_j = (|\bar{Y}_{T_j} - y_0| + 1)^2$, $V_j = \sup_{t \in [T_j, T_{j+1} \wedge 1)} |W_t - W_{T_j}|^2$ and $\mathcal{G}_j = \sigma(W_t : t \leq T_j)$ to conclude that

$$\mathbb{E}\Big[\sup_{t\in[0,1]}|A_t|^2\Big] \le K^2|\Sigma|^2 \mathbb{E}\Big[\sup_{j=0,\dots,r-1}(|\bar{Y}_{T_j}-y_0|+1)^2\Big] \cdot \mathbb{E}\Big[\sup_{j=0,\dots,r-1}\sup_{t\in[T_j,T_{j+1}\wedge 1)}|W_t-W_{T_j}|^2\Big].$$

By Lévy's modulus of continuity, the term

$$||W||_{\varphi} := \sup_{0 \le s < t \le 1} \frac{|W_t - W_s|}{\varphi(t-s)}$$

is almost surely finite for $\varphi : [0,1] \to [0,\infty), \delta \mapsto \sqrt{\delta \log(e/\delta)}$. Hence we get (for instance with the isoperimetric inequality) that $\mathbb{E} \|W\|_{\varphi}^2$ is finite. Recalling that neighboring points in \mathbb{T} are at most ε units apart, we conclude with the monotonicity of φ on [0,1] that

$$\mathbb{E}\left[\sup_{j=0,\dots,r-1}\sup_{t\in[T_j,T_{j+1}\wedge 1)}|W_t-W_{T_j}|^2\right] \leq \mathbb{E}\left[\|W\|_{\varphi}^2\right]\varphi(\varepsilon)^2$$

so that

$$\mathbb{E}\Big[\sup_{t\in[0,1]}|A_t|^2\Big] \le K^2|\Sigma|^2 \mathbb{E}\Big[\sup_{j=0,\dots,r-1}(|\bar{Y}_{T_j}-y_0|+1)^2\Big] \mathbb{E}\Big[\|W\|_{\varphi}^2\Big] \varphi(\varepsilon)^2.$$
(17)

In the general case (L'_t) is a Lévy process that is independent of the Wiener process (W_t) . Disintegrating the Lévy process allows us to apply the above result and to retrieve (17) also for random processes (L'_t) . Combining (16) and (17), we get

$$\mathbb{E}[\|\bar{Y} - \hat{Y}\|^2] \le 2K^2 \,\mathbb{E}[(\|\bar{Y} - y_0\| + 1)^2] (|\Sigma|^2 \,\mathbb{E}[\|W\|_{\varphi}^2] \,\varphi(\varepsilon)^2 + |b - F_0(h)|^2 \varepsilon^2).$$

Next, note that by, Proposition 1 and Lemma 1, $\mathbb{E}[(\|\bar{Y} - y_0\| + 1)^2]$ is bounded from above by some constant depending on K only. Consequently, there exists a constant κ with

$$\mathbb{E}[\|\bar{Y} - \hat{Y}\|^2] \le \kappa (|\Sigma|^2 \varphi(\varepsilon)^2 + |b - F_0(h)|^2 \varepsilon^2).$$

Together with Proposition 1, one immediately obtains the statement for the case without Wiener process. In order to obtain the statement of the general case, we again use that $|\int_{B_h^c} x \nu(dx)|^2 \leq K^2/\varepsilon$ due to the Cauchy-Schwarz inequality.

Proof of part (i) of Theorem 1

We can assume without loss of generality that

$$\frac{1}{h^{2/3}} \precsim g(h) \precsim \frac{1}{h(\log 1/h)^{1+\gamma}},\tag{18}$$

since otherwise, we can modify g in such a way that the new g is larger than the old one and enjoys the wanted properties. We consider a multilevel Monte Carlo algorithm \widehat{S} (as introduced in Section 4) with $h_k = g^{-1}(2^k)$ and $\varepsilon_k = 2^{-k}$ for $k \in \mathbb{Z}_+$. For technical reasons, we also define ε_0 and h_0 although they do not appear in the algorithm. The parameters $m \in \mathbb{N}$ and $n_1, \ldots, n_m \in \mathbb{N}$ are specified below. Note that

$$\nu(B_{h_k}^c) \le g(h_k) = 1/\varepsilon_k \quad \text{and} \quad \varepsilon_k \le 1, \tag{19}$$

so that by Theorem 2, we have

$$\begin{split} \mathbb{E}[\|Y^{(k)} - Y^{(k-1)}\|^2] &\leq 2\mathbb{E}[\|Y - Y^{(k)}\|^2] + 2\mathbb{E}[\|Y - Y^{(k-1)}\|^2] \\ &\leq \kappa_1 \Big[F(h_{k-1}) + |b - F_0(h_k)|^2 \varepsilon_{k-1}^2\Big]. \end{split}$$

for some constant $\kappa_1 > 0$. By Lemma 1 and Theorem 2, $\mathbb{E}[||Y^{(1)} - y_0||^2]$ is bounded from above by some constant depending on K only. Hence, equation (10) together with Theorem 2 imply the existence of a constant κ_2 such that

$$e^{2}(\widehat{S}) \leq \kappa_{2} \sum_{k=1}^{m+1} \frac{1}{n_{k}} \left[F(h_{k-1}) + |b - F_{0}(h_{k})|^{2} \varepsilon_{k-1}^{2} \right],$$
(20)

where we set $n_{m+1} = 1$. Next, we analyze the terms in (20). Using assumption (3), we have, for a sufficiently small $v \in (0, 1)$ and an appropriate constant κ_3 ,

$$|F_0(h_k)| \le \int |x| \,\nu(\mathrm{d}x) \le \frac{1}{v} \int |x| (v \lor |x|) \,\nu(\mathrm{d}x) \le \frac{1}{v} \int_{B_v^c} |x|^2 \,\nu(\mathrm{d}x) + \int_0^v \nu(B_u^c) \,\mathrm{d}u$$
$$\le \frac{1}{v} \int |x|^2 \,\nu(\mathrm{d}x) + \kappa_3 \int_0^v \frac{1}{u(\log 1/u)^{1+\gamma}} \,\mathrm{d}u.$$

Both integrals are finite. Moreover, we have

$$F(h_k) \le g(h_k) h_k^2 = 2^k g^{-1} (2^k)^2, \qquad (21)$$

and using (18) we get that

$$\frac{1}{y^{3/2}} \precsim g^{-1}(y) \precsim \frac{1}{y(\log y)^{1+\gamma}} \qquad \text{as } y \to \infty.$$
(22)

Hence, we have $2^k g^{-1} (2^k)^2 \gtrsim 2^{-2k} = \varepsilon_k^2$ as k tends to infinity, and there exists a constant κ_4 such that

$$e^{2}(\widehat{S}) \leq \kappa_{4} \sum_{k=1}^{m+1} \frac{1}{n_{k}} 2^{k-1} g^{-1} (2^{k-1})^{2}.$$

We shall now fix m and n_1, \ldots, n_m . For a given parameter $Z \ge 1/g^{-1}(1)$, we choose $m = m(Z) = \inf\{k \in \mathbb{N} : Zg^{-1}(2^k) < 1\} - 1$. Moreover, we set $n_k = n_k(Z) = \lfloor Zg^{-1}(2^{k-1}) \rfloor$ for $k = 1, \ldots, m$, and set again $n_{m+1} = 1$. Then $1/n_k \le 2/(Zg^{-1}(2^{k-1}))$ for $k = 1, \ldots, m+1$, so that

$$e^{2}(\widehat{S}) \leq \kappa_{4} \sum_{k=1}^{m+1} \frac{1}{n_{k}} 2^{k-1} g^{-1} (2^{k-1})^{2} \leq 2\kappa_{4} \frac{1}{Z} \sum_{k=1}^{m+1} 2^{k-1} g^{-1} (2^{k-1}).$$

By (22), $2^k g^{-1}(2^k) \preceq k^{-(1+\gamma)}$ and the latter sum is uniformly bounded for all m. Hence, there exists a constant κ_5 depending only on g and K such that

$$e^2(\widehat{S}) \le \kappa_5 \frac{1}{Z}.$$

It remains to analyze the cost of the algorithm. The expected number of breakpoints of $Y^{(k)}$ is less than $1/\varepsilon_k + \nu(B_{h_k}^c) \leq 2^{k+1}$ (see (19)) so that

$$\operatorname{cost}(\widehat{S}) \le \sum_{k=1}^{m} 2^{k+1} n_k.$$
(23)

Hence,

$$\operatorname{cost}(\widehat{S}) \le 4Z \sum_{k=1}^{m} 2^{k-1} g^{-1}(2^{k-1}) \le \kappa_6 Z_5$$

where $\kappa_6 > 0$ is an appropriate constant.

Proof of part (ii) of Theorem 1

We proceed similarly as in the proof of part (i). We assume without loss of generality that g satisfies

$$\frac{\sqrt{\log 1/h}}{h} \precsim g(h) \precsim \frac{(\log 1/h)^{\gamma}}{h},\tag{24}$$

since, otherwise, we can enlarge g appropriately. In analogy to the proof of part (i), we choose $h_k = g^{-1}(2^k)$ and $\varepsilon_k = 2^{-k}$ for $k \in \mathbb{Z}_+$, and we note that estimates (19) and (21) remain valid. Next, we deduce with equation (10) and Theorem 2 that, for some constant κ_1 ,

$$e^{2}(\widehat{S}) \leq \kappa_{1} \sum_{k=1}^{m+1} \frac{1}{n_{k}} \Big[F(h_{k-1}) + \varepsilon_{k-1} \log \frac{e}{\varepsilon_{k-1}} \Big] \leq \kappa_{1} \sum_{k=1}^{m+1} \frac{1}{n_{k}} \Big[2^{k-1} g^{-1} (2^{k-1})^{2} + 2^{-(k-1)} \log(e2^{k-1}) \Big],$$

where again $n_{m+1} = 1$. Note that (24) implies that

$$\frac{\sqrt{\log y}}{y} \precsim g^{-1}(y) \precsim \frac{(\log y)^{\gamma}}{y} \quad \text{as } y \to \infty,$$
(25)

so that, in particular, $2^k g^{-1} (2^k)^2 \succeq 2^{-k} \log(e2^k)$ as k tends to infinity. Hence, we find a constant κ_2 such that

$$e^{2}(\widehat{S}) \leq \kappa_{2} \sum_{k=1}^{m+1} \frac{1}{n_{k}} 2^{k-1} g^{-1} (2^{k-1})^{2}.$$

For a parameter $Z \ge e \lor (1/g^{-1}(1))$, we choose $m = m(Z) = \inf\{k \in \mathbb{N} : Zg^{-1}(2^k) < 1\} - 1$, and we set $n_k = n_k(Z) = \lfloor Zg^{-1}(2^{k-1}) \rfloor$ for k = 1, ..., m. Then we get with (25) that

$$e^{2}(\widehat{S}) \leq 2\kappa_{2} \frac{1}{Z} \sum_{k=1}^{m+1} 2^{k-1} g^{-1}(2^{k-1}) \leq \kappa_{3} \frac{1}{Z} m^{\gamma+1}$$

for an appropriate constant κ_3 . By definition, $Zg^{-1}(2^m) \ge 1$ so that, by equation (24),

$$m \le \log g(\frac{1}{Z}) / \log 2 \precsim \log Z$$
 as $Z \to \infty$.

Consequently, there exists a constant κ_4 such that

$$e^2(\widehat{S}) \le \kappa_4 \frac{(\log Z)^{\gamma+1}}{Z}.$$

Similarly, we get for the cost function

$$\operatorname{cost}(\widehat{S}) \le \sum_{k=1}^{m} 2^{k+1} n_k \le 4Z \sum_{k=1}^{m} 2^{k-1} g^{-1} (2^{k-1}) \le \kappa_5 Z (\log Z)^{\gamma+1}.$$

Next, we choose

$$Z = Z(n) = \frac{1}{2\kappa_5} \frac{n}{(\log n)^{\gamma+1}}$$

for $n \ge e$ sufficiently large such that $Z \ge e \lor (1/g^{-1}(1))$. Then

$$cost(\widehat{S}) \le \kappa_5 Z (\log Z)^{\gamma+1} \sim \frac{1}{2}n$$

and we have, for all sufficiently large n, $\operatorname{cost}(\widehat{S}) \leq n$. Conversely, we find

$$e^2(\widehat{S}) \le \kappa_4 \frac{(\log Z)^{\gamma+1}}{Z} \approx \frac{(\log n)^{2(\gamma+1)}}{n}.$$

Proof of part (iii) of Theorem 1

First note that property (4) is equivalent to

$$\frac{\gamma}{2}g^{-1}(u) \le g^{-1}(2u) \tag{26}$$

for all sufficiently large u > 0. This implies that there exists a finite constant κ_1 depending only on g such that for all $k, l \in \mathbb{Z}_+$ with $k \leq l$ one has

$$g^{-1}(2^k) \le \kappa_1 \left(\frac{2}{\gamma}\right)^{l-k} g^{-1}(2^l).$$
(27)

We proceed similar as in the proof of part (i) and consider $\hat{S} \in \mathcal{A}$ with $h_k = g^{-1}(2^k)$ and $\varepsilon_k = 2^{-k}$ for $k \in \mathbb{Z}_+$. The maximal index m and the number of iterations n_k are fixed later in the discussion. Again estimates (19) and (21) remain valid and we get with Theorem 2

$$e^{2}(\widehat{S}) \leq \kappa_{2} \sum_{k=1}^{m+1} \frac{1}{n_{k}} \Big[F(h_{k-1}) + \varepsilon_{k-1} \log \frac{e}{\varepsilon_{k-1}} \Big] \leq \kappa_{2} \sum_{k=1}^{m+1} \frac{1}{n_{k}} \Big[g^{-1} (2^{k-1})^{2} 2^{k-1} + 2^{-(k-1)} \log(e2^{k-1}) \Big],$$

for a constant κ_2 and $n_{m+1} = 1$ as before. By (26), we have $g^{-1}(2^k) \succeq (\gamma/2)^k$ and recalling that $\gamma > 1$ we conclude that $2^{-k} \log(e2^k) \preceq g^{-1}(2^k)^2 2^k$ as k tends to infinity. Hence, there exists a constant κ_3 with

$$e^{2}(\widehat{S}) \leq \kappa_{3} \sum_{k=1}^{m+1} \frac{1}{n_{k}} 2^{k-1} g^{-1} (2^{k-1})^{2}.$$

Conversely, we again estimate the cost by

$$\operatorname{cost}(\widehat{S}) \le \sum_{k=1}^{m} 2^{k+1} n_k.$$

Now we specify m and n_1, \ldots, n_m . For a given parameter $Z \ge 2/g^{-1}(1)$, we let $m = m(Z) = \inf\{k \in \mathbb{N} : Zg^{-1}(2^k) < 2\} - 1$, and set $n_k = n_k(Z) = \lfloor Zg^{-1}(2^{k-1}) \rfloor$ for $k = 1, \ldots, m$. Then we get with (27) that there exists a constant κ_4 with

$$e^{2}(\widehat{S}) \leq 2\kappa_{3}\frac{1}{Z}\sum_{k=1}^{m+1} 2^{k-1}g^{-1}(2^{k-1}) \leq 2\kappa_{1}\kappa_{3}\frac{1}{Z}\sum_{k=1}^{m+1} 2^{k-1}g^{-1}(2^{m+1})\left(\frac{2}{\gamma}\right)^{m+1-(k-1)}$$
$$\leq 2\kappa_{1}\kappa_{3}\frac{1}{Z}2^{m+1}g^{-1}(2^{m+1})\sum_{k=1}^{m+1}\gamma^{-(m+1-(k-1))} \leq \kappa_{4}\frac{1}{Z}2^{m+1}g^{-1}(2^{m+1}).$$

Moreover, by (26) one has for sufficiently large m (or, equivalently, for sufficiently large Z) that $Zg^{-1}(2^{m+1}) \geq \frac{\gamma}{2}Zg^{-1}(2^m) \geq \gamma > 1$ so that

$$e^2(\widehat{S}) \le \kappa_4 2^{m+1} g^{-1} (2^{m+1})^2$$

Similarly, one obtains

$$cost(\widehat{S}) \le \kappa_5 Z 2^{m+1} g^{-1}(2^{m+1}) < 2\kappa_5 2^{m+1}$$

Next, we choose, for given $n \ge 2\kappa_5$, Z > 0 such that $m = \lfloor \log_2 \frac{n}{2\kappa_5} \rfloor - 1$. Then, clearly, $\cos(\widehat{S}) \le n$ and for sufficiently large n we have

$$e^2(\widehat{S}) \le \kappa_4 \frac{n}{2\kappa_5} g^{-1} (\frac{n}{4\kappa_5})^2 \precsim n g^{-1}(n)^2 \text{ as } n \to \infty.$$

In the last step, we again used property (26).

Proof of Corollary 1

We assume without loss of generality that $\beta < 2$ since otherwise the statement of the corollary is trivial. We fix $\beta' \in (\beta, 2]$ and recall that

$$C := \int_{B_1} |x|^{\beta'} \,\nu(\mathrm{d}x)$$

is finite. We consider $\bar{g}: (0,\infty) \to (0,\infty), h \mapsto \int \frac{|x|^2}{h^2} \wedge 1\nu(\mathrm{d}x)$, whose integral we split for $h \in (0,1]$ into three parts:

$$\bar{g}(h) = \int_{B_h} \frac{|x|^2}{h^2} \nu(\mathrm{d}x) + \int_{B_1 \setminus B_h} 1 \nu(\mathrm{d}x) + \int_{B_1^c} 1 \nu(\mathrm{d}x) =: I_1 + I_2 + I_3.$$

The last term does not depend on h and we estimate the first two terms by

$$I_1 \le h^{-\beta'} \int_{B_h} |x|^{\beta'} \nu(\mathrm{d}x) \le C h^{-\beta'} \text{ and } I_2 \le h^{-\beta'} \int_{B_1 \setminus B_h} |x|^{\beta'} \nu(\mathrm{d}x) \le C h^{-\beta'}.$$

Hence, we can choose $\beta'' \in ((\beta' \vee 1), 2]$ arbitrarily, and a decreasing and invertible function $g: (0, \infty) \to (0, \infty)$ that dominates \bar{g} and satisfies $g(h) = h^{-\beta''}$ for all sufficiently small h > 0. Then for $\gamma = 2^{1-1/\beta''}$, one has $g(\frac{\gamma}{2}h) = 2g(h)$ for all sufficiently small h > 0 and we are in the position to apply Part (iii) of Theorem 1 to get

$$e_{\mathcal{A}}(n) \precsim n^{\frac{1}{2} - \frac{1}{\beta''}}.$$

In the case where $\beta < 1$, one can choose $\beta' = 1$ and $\beta'' > 1$ arbitrarily close to one and gets the result. Whereas for $\beta \ge 1$, one can choose for any $\beta'' > \beta$ an appropriate β' and thus retrieve the statement.

Appendix

We shall use the following consequence of the Itô isometry for Lévy processes.

Lemma 3. Let (A_t) be a previsible process with state space $\mathbb{R}^{d_Y \times d_X}$, let (L_t) be a square integrable \mathbb{R}^{d_X} -valued Lévy martingale and denote by $\langle L \rangle$ the process given via

$$\langle L \rangle_t = \sum_{j=1}^{d_X} \langle L^{(j)} \rangle_t,$$

where $\langle L^{(j)} \rangle$ denotes the predictable compensator of the classical bracket process for the *j*-th coordinate of *L*. One has, for any stopping time τ with finite expectation $\mathbb{E} \int_0^{\tau} |A_s|^2 d\langle L \rangle_s$, that $(\int_0^{t\wedge\tau} A_s dL_s)_{t\geq 0}$ is a uniformly square integrable martingale which satisfies

$$\mathbb{E} \left| \int_0^\tau A_s \, \mathrm{d}L_s \right|^2 \le \mathbb{E} \int_0^\tau |A_s|^2 \, \mathrm{d}\langle L \rangle_s.$$

In general the estimate can be strengthened by replacing the Frobenius norm by the matrix norm induced by the Euclidean norm. For the convenience of the reader we provide a proof of the lemma. *Proof.* Let ν denote the Lévy measure of L and $\Sigma\Sigma^*$ the covariance of its Wiener component. Let $Q: \mathbb{R}^{d_X} \to \mathbb{R}^{d_X}$ denote the self-adjoint operator given by

$$Qx = \Sigma \Sigma^* x + \int \langle x, y \rangle \, y \, \nu(\mathrm{d}y).$$

We recall the Itô isometry for Lévy processes. One has for a previsible process (A_s) with

$$\mathbb{E}\int_0^\tau |A_s Q^{1/2}|^2 \,\mathrm{d}s < \infty,$$

that $(\int_0^{t\wedge\tau} A_s \, \mathrm{d}L_s)_{t\geq 0}$ is a uniformly square integrable martingale with

$$\mathbb{E}\left|\int_0^\tau A_s \,\mathrm{d}L_s\right|^2 = \mathbb{E}\int_0^\tau |A_s Q^{1/2}|^2 \,\mathrm{d}s.$$

The statement now follows immediately by noticing that

$$|A_s Q^{1/2}|^2 \le |A_s|^2 \operatorname{tr}(Q) \text{ and } \int_0^\tau |A_s|^2 \operatorname{tr}(Q) \, \mathrm{d}s = \int_0^\tau |A_s|^2 \, \mathrm{d}\langle L \rangle_s.$$

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